

ON THE OPTIMALITY OF THE MONTE-CARLO ESTIMATOR

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ABSTRACT. We prove that on an atomless probability space, the worst-case mean squared error of the Monte-Carlo estimator is minimal if the random points are chosen independently.

1. INTRODUCTION AND STATEMENT OF THE RESULTS

Let (X, μ) be a probability space. We are interested in the following general question: if f is a measurable, real or complex-valued function on X , how can we efficiently compute the integral $\int_X f \, d\mu$? The famous *Monte-Carlo method* is a solution to this problem: just choose an integer n big enough, and draw Z_1, \dots, Z_n independent X -valued random variables (that is, *random points*) of law μ , and form the mean $\frac{1}{n} \sum_{i=1}^n f(Z_i)$, called the *Monte-Carlo estimator*.

We measure the quality of this method by computing what we call the *mean squared error*: we have the well-known equality, valid for all $n \in \mathbb{N}^*$ and $f \in L^2(X, \mu)$,

$$\text{Var} \left(\frac{1}{n} \sum_{i=1}^n f(Z_i) - \int_X f \, d\mu \right) = \frac{1}{n} \left\| f - \int_X f \, d\mu \right\|_{L^2(X, \mu)}^2$$

and we obtain the following equality, concerning the *worst-case mean squared error*:

$$\sup_{\substack{f \in L^2(X, \mu) \\ \|f\|_2=1}} \text{Var} \left(\frac{1}{n} \sum_{i=1}^n f(Z_i) - \int_X f \, d\mu \right) = \frac{1}{n}.$$

In this paper, we study the question of measuring the worst-case mean squared error, in the general situation where the points Z_i are not supposed independent, and we prove the following theorem and its corollary.

Theorem. *Let (X, μ) be a probability space, let $N, n \in \mathbb{N}^*$, and $Z := (Z_1, \dots, Z_n)$ an n -tuple of random points on X such that for all i , the law of Z_i is μ . We do not assume that the Z_i 's are independent. Furthermore, we assume that X can be partitioned in N measurable subsets of equal measure.*

We then have

$$\sup_{\substack{f \in L^2(X, \mu) \\ \|f\|_2=1}} \text{Var} \left(\frac{1}{n} \sum_{i=1}^n f(Z_i) - \int_X f \, d\mu \right) \geq \frac{1}{n} \left(1 - \frac{n-1}{N-1} \right).$$

Corollary. *Let (X, μ) be an atomless probability space, $n \in \mathbb{N}^*$, and $Z := (Z_1, \dots, Z_n)$ an n -tuple of random points on X such that for all i , the law of Z_i is μ . We do not assume that the Z_i 's are independent.*

We then have

$$\sup_{\substack{f \in L^2(X, \mu) \\ \|f\|_2=1}} \text{Var} \left(\frac{1}{n} \sum_{i=1}^n f(Z_i) - \int_X f \, d\mu \right) \geq \frac{1}{n}.$$

Remark. *As we shall see in the paper, in the case where $X := \{1, \dots, N\}$ and μ is the uniform measure on X , the inequality of the theorem is an equality when the law of Z is the uniform measure on the set of n -tuples of points in X such that the coordinates are pairwise different. This random n -tuple is then, in the sense of the worst-case mean squared error, than an independent n -tuple.*

As we saw before, the inequality in the corollary is an equality if the Z_i 's are independent. We don't know if this condition is necessary. It is, to our opinion, worth knowing that in [LPS86], the authors build, for all prime p such that $p \equiv 1[4]$, a $(p+1)$ -tuple Z of uniform random points on the 2-sphere which are not independent, and prove that its worst-case mean squared error is $\frac{4p}{(p+1)^2}$, which is approximately 4 times the lower bound in the corollary. In the article [LP18], it is shown that their construction is optimal, in a broad framework.

We confess our astonishment of having found no trace of these statements, which answer a question that we find both natural and general, and in an elementary way.

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2. PROOFS

To alleviate the presentation, we use the following notation: we consider the numbers

$$\text{MSE}_Z(f) := \text{Var} \left(\frac{1}{n} \sum_{i=1}^n f(Z_i) - \int_X f \, d\mu \right)$$

et

$$\text{MSE}(Z) := \sup_{\substack{f \in L^2(X, \mu) \\ \|f\|_2=1}} \text{MSE}_Z(f).$$

First of all, if $f \in L^2(X, \mu)$, we notice that $\text{MSE}_Z(f) = \text{MSE}(Z) (f - \int_X f \, d\mu)$. Consequently, $\text{MSE}(Z)$ is also the sup of the $\text{MSE}_Z(f)$ for f of norm 1 zero integral.

Let $f \in L^2(X, \mu)$, of norm 1 and zero integral. We have that

$$\begin{aligned} \text{MSE}_Z(f) &= \mathbb{E} \left[\left(\frac{1}{n} \sum_{i=1}^n f(Z_i) \right)^2 \right] \\ &= \mathbb{E} \left[\frac{1}{n^2} \sum_{i=1}^n f(Z_i)^2 + \frac{1}{n^2} \sum_{i \neq j} f(Z_i) f(Z_j) \right] \\ &= \frac{1}{n^2} \sum_{i=1}^n \mathbb{E}[f(Z_i)^2] + \frac{1}{n^2} \sum_{i \neq j} \mathbb{E}[f(Z_i) f(Z_j)] \\ &= \frac{1}{n} + \frac{1}{n^2} \sum_{i \neq j} \mathbb{E}[f(Z_i) f(Z_j)] \end{aligned}$$

and we recover the fact recalled above: if the Z_i 's are pairwise independent, and if f is of norm 1 of zero integral, $\text{MSE}_Z(f) = \frac{1}{n}$.

Let us prove the theorem.

Proof of the theorem. Let X_1, \dots, X_N be measurable subsets that partition X , all of measure $\frac{1}{N}$, with $N \geq 2$. Let us denote, for $p \in \{1, \dots, N\}$, $\mu_p := \mu(X_p)$. For every $(p, q) \in \{1, \dots, N\}^2$, we set

$$f_{p,q} := \sqrt{\frac{N}{2}} \mathbf{1}_{X_p} - \sqrt{\frac{N}{2}} \mathbf{1}_{X_q}.$$

Moreover, we will denote, for $k \in \{1, \dots, N\}$, $f_{p,q}(X_k)$ the value that $f_{p,q}$ takes on X_k - this abuse of notation is harmless because $f_{p,q}$ is constant on the X_i 's.

$f_{p,q}$ is visibly of zero integral, and if $p \neq q$, its norm is 1.

We will prove that there are different $p, q \in \{1, \dots, N\}$ such that $\text{MSE}_Z(f_{p,q}) \geq \frac{1}{n} \left(1 - \frac{n-1}{N-1}\right)$.

Let $p, q \in \{1, \dots, N\}$. We have that

$$\begin{aligned}
\text{MSE}_Z(f_{p,q}) &= \frac{1}{n} + \frac{1}{n^2} \sum_{i \neq j} \mathbb{E}[f_{p,q}(Z_i)f_{p,q}(Z_j)] \\
&= \frac{1}{n} + \frac{1}{n^2} \sum_{i \neq j} \left(\sum_k \mathbb{P}(Z_i \in X_k \text{ et } Z_j \in X_k) f_{p,q}(X_k)^2 \right. \\
&\quad \left. + \sum_{l \neq m} \mathbb{P}(Z_i \in X_l \text{ et } Z_j \in X_m) f_{p,q}(X_l) f_{p,q}(X_m) \right) \\
&= \frac{1}{n} + \frac{1}{n^2} \frac{N}{2} \sum_{i \neq j} (\mathbb{P}(Z_i \in X_p \text{ et } Z_j \in X_p) \\
&\quad + \mathbb{P}(Z_i \in X_q \text{ et } Z_j \in X_q) \\
&\quad - \mathbb{P}(Z_i \in X_p \text{ et } Z_j \in X_q) \\
&\quad - \mathbb{P}(Z_i \in X_q \text{ et } Z_j \in X_p)).
\end{aligned}$$

from which we deduce the inequality

$$\text{MSE}_Z(f_{p,q}) \geq \frac{1}{n} - \frac{1}{n^2} \frac{N}{2} \left(\sum_{i \neq j} \mathbb{P}(Z_i \in X_p \text{ et } Z_j \in X_q) + \mathbb{P}(Z_i \in X_q \text{ et } Z_j \in X_p) \right).$$

Let us denote

$$\theta_{p,q} := \sum_{i \neq j} \mathbb{P}(Z_i \in X_p \text{ et } Z_j \in X_q) + \mathbb{P}(Z_i \in X_q \text{ et } Z_j \in X_p).$$

Let us compute:

$$\begin{aligned}
\sum_{p \neq q} \theta_{p,q} &= 2 \sum_{p \neq q} \sum_{i \neq j} \mathbb{P}(Z_i \in X_p \text{ et } Z_j \in X_q) \\
&= 2 \sum_{i \neq j} \sum_{p \neq q} \mathbb{P}(Z_i \in X_p \text{ et } Z_j \in X_q) \\
&= 2 \sum_{i \neq j} \mathbb{P}(Z_i \text{ et } Z_j \text{ ne sont pas dans le même morceau de la partition}) \\
&\leq 2n(n-1).
\end{aligned}$$

Now, since this sum of $N(N-1)$ numbers is lower or equal than $2n(n-1)$, then one of the terms must be lower or equal than $2 \frac{n(n-1)}{N(N-1)}$. For a couple (p, q) such that $\theta_{p,q} \leq 2 \frac{n(n-1)}{N(N-1)}$, we then have

$$\begin{aligned}
\text{MSE}_Z(f_{p,q}) &\geq \frac{1}{n} - \frac{1}{n^2} \frac{N}{2} 2 \frac{n(n-1)}{N(N-1)} \\
&= \frac{1}{n} - \frac{n-1}{n(N-1)} \\
&= \frac{1}{n} \left(1 - \frac{n-1}{N-1} \right).
\end{aligned}$$

□

Here's an example where the inequality is an equality.

Proposition. *If $X := \{1, \dots, N\}$, if μ is the uniform probability on X , if $n \leq N$, and if the law of Z is the uniform measure on the set of n -tuples of points in X which coordinates are pairwise different, then the inequality in the theorem is an equality, that is,*

$$\text{MSE}_Z = \frac{1}{n} \left(1 - \frac{n-1}{N-1} \right).$$

Proof. Let π be the measure on X^n defined by

$$\pi := \frac{(N-n)!}{N!} \sum_{\substack{i_1, \dots, i_n \in X \\ \forall j \neq k, \\ i_j \neq i_k}} \delta_{(i_1, \dots, i_n)},$$

where δ is the notation for a Dirac measure. In words, π is the uniform measure on the set of n -tuples of points in X which coordinates are pairwise different. Let $Z = (Z_1, \dots, Z_n)$ be an n -tuple of law π (we then have, for all i , that Z_i is uniform on X).

Let $f \in L^2(X, \mu)$ be of norm 1, and such that $\int f d\mu = 0$. Let us compute:

$$\begin{aligned}
\sum_{l \neq m} \mathbb{E}[f(Z_l)f(Z_m)] &= \sum_{l \neq m} \mathbb{E} \left[\sum_{\substack{A \subset X \\ |A|=n}} \sum_{\substack{i_1, \dots, i_n \in A \\ \forall j \neq k \\ i_j \neq i_k}} \mathbf{1}_{\{Z_1=i_1, \dots, Z_n=i_n\}} f(Z_l)f(Z_m) \right] \\
&= \sum_{l \neq m} \sum_{\substack{A \subset X \\ |A|=n}} \sum_{\substack{i_1, \dots, i_n \in A \\ \forall j \neq k \\ i_j \neq i_k}} \mathbb{P}[Z_1 = i_1, \dots, Z_n = i_n] f(i_l)f(i_m) \\
&= \binom{N}{n}^{-1} \sum_{\substack{A \subset X \\ |A|=n}} \sum_{\substack{p, q \in A \\ p \neq q}} f(p)f(q) \\
&= \binom{N-2}{n-2} \binom{N}{n}^{-1} \sum_{\substack{p, q \in X \\ p \neq q}} f(p)f(q) \\
&= \frac{n(n-1)}{N(N-1)} \sum_{p \in X} f(p) \sum_{\substack{q \in X \\ q \neq p}} f(q) \\
&= \frac{n(n-1)}{N-1} \|f\|_2^2 \\
&= \frac{n(n-1)}{N-1}.
\end{aligned}$$

We therefore have

$$\text{MSE}_Z(f) = \frac{1}{n} \left(1 - \frac{n-1}{N-1} \right).$$

□

Let us prove the corollary.

Proof of the corollary. We will prove that for all $\epsilon > 0$, we have that $\text{MSE}(Z) \geq \frac{1}{n} - \epsilon$, which is enough. According to a theorem of Sierpiński [Sie22], every atomless probability space is such that for every $a \in [0, 1]$, there is a measurable subset of X of measure a . From this, it is easy, for all arbitrarily big N , to partition X in N of measurable subsets of equal measure. If we choose N such that $\frac{1}{n} \left(1 - \frac{n-1}{N-1} \right) \geq \frac{1}{n} - \epsilon$, which is obviously possible, then according to the theorem, it is possible to find f of norm 1, zero integral, such that $\text{MSE}_Z(f) \geq \frac{1}{n} - \epsilon$. □

For the sake of completeness, we add a simple proof of Sierpiński's theorem.

Complement (Sierpiński's theorem on atomless probability spaces). *If (X, \mathcal{B}, μ) is an atomless probability space, then for every measurable $A \subset X$, there exists $\phi : [0, \mu(A)] \rightarrow \mathcal{B}$ non-decreasing, such that $\forall t \in [0, \mu(A)], \mu(\phi(t)) = t$.*

Proof. The hypothesis of X being atomless means that for every measurable $B \subset X$ such that $\mu(B) > 0$, there exists a measurable $C \subset B$ such that $0 < \mu(C) < \mu(B)$.

Let A be a measurable subset of X , such that $\mu(A) > 0$ (if $\mu(A) = 0$, it is enough to define $\phi(0) := A$). By applying Zorn's lemma, we obtain a $\phi : I \rightarrow \mathcal{B}$ where I is a subset of $[0, \mu(A)]$, ϕ is non-decreasing, such that $\forall i \in I, \mu(\phi(i)) = i$, and such that ϕ has no strict extension that satisfies these properties. Let us show that I equals $[0, \mu(A)]$.

On the one hand, I is closed. Indeed, let $(x_n)_{n \in \mathbb{N}}$ be a sequence of elements in I that converges to some x . Let us show that $x \in I$. We can assume, up to extracting a subsequence, that $(x_n)_n$ is monotonous. If $x \notin I$, let us define $\tilde{\phi} := I \cup \{x\} \rightarrow \mathcal{B}$ that extends ϕ by defining

$\tilde{\phi}(x) := \bigcap_n \phi(x_n)$ if $(x_n)_n$ is non-increasing, and $\tilde{\phi}(x) := \bigcup_n \phi(x_n)$ if $(x_n)_n$ non-decreasing. According to μ 's continuity properties, $\mu(\phi(x)) = \lim_n x_n = x$, and according to the monotony properties of μ , $\tilde{\phi}$ is non-decreasing. $\tilde{\phi}$ is therefore a strict extension of ϕ that verifies the same properties. This is a contradiction. So $x \in I$, and therefore, I is closed.

On the other hand, I verifies $\forall a, b \in I, a < b \Rightarrow (\exists c \in I, a < c < b)$ (we say that I is *order-dense*). Indeed, if there are $a, b \in I$ such that $a < b$ and $]a, b[\cap I = \emptyset$, then let us use the hypothesis that X is atomless, which provides a measurable $C \subset \phi(b) \setminus \phi(a)$ such that $0 < \mu(C) < b - a$. Let us then define $\tilde{\phi} : I \cup \{a + \mu(C)\}$ that extends ϕ by defining $\tilde{\phi}(a + \mu(C)) := \phi(a) \cup C$. Then $\mu(\tilde{\phi}(a + \mu(C))) = \mu(\phi(a) \cup C) = a + \mu(C)$. According to μ 's monotony properties, $\tilde{\phi}$ est non-decreasing. $\tilde{\phi}$ is then a strict extension of ϕ that verifies the same properties. This is a contradiction. Therefore, I is order-dense.

So I is closed and order-dense. Therefore, $I = [0, \mu(A)]$. □

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